Instrumental Variable Models for Discrete Outcomes

Department Seminar: UIUC Economics Department

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CeMMAP & UCL

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IV Models for Discrete Outcomes

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Y = h(X, U)

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 - Extensions/applications.

• $Y \in \{0, 1, ..., M\}$ determined by X and $U \sim Unif(0, 1)$: $Y = h(X, U) \qquad h \uparrow U \qquad U \perp Z$

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• $Y \in \{0, 1, \dots, M\}$ determined by X and $U \sim Unif(0, 1)$:

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• Threshold crossing representation. Consider some h_0 .

$$h_0(x, u) = \begin{cases} 0 & , & 0 < u \le p_0^0(x) \\ 1 & , & p_0^0(x) < u \le p_1^0(x) \\ \vdots & \vdots & \vdots & \vdots \\ M & , & p_{M-1}^0(x) < u \le 1 \end{cases}$$

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$$F_{UX|Z}^{0}(u,x|z) \equiv \Pr[U \le u \cap X \le x|Z=z]$$

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• It determines a distribution function of Y and X given Z

$$F_{YX|Z}^{0}(m, x|z) = F_{UX|Z}^{0}(p_{m}^{0}(x), x|z)$$

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• $U \perp Z$ limits adjustment of the U and X arguments of admissible $F_{UX|Z}$ because for all τ , z

$$F_{UX|Z}(\tau, \infty|z) \equiv F_{U|Z}(\tau|z) = F_U(\tau) = \tau$$

Some related results:

• **Continuous outcomes**: Chernozhukov and Hansen (2005) and related papers.

Y = h(X, U) $U \perp Z$ h strictly increasing

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• Triangular models: structural equation for (continuous) X:

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Chesher (2003, 2005), Imbens & Newey (2003).

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Chesher (2003, 2005), Imbens & Newey (2003).

• Simultaneous models: "single equation" analysis of Tamer's (2003) entry game.

$$Y_{1}^{*} = \alpha_{1} Y_{2} + Z\beta_{1} + \varepsilon_{1} \qquad Y_{2}^{*} = \alpha_{2} Y_{1} + Z\beta_{2} + \varepsilon_{2}$$
$$Y_{1} = \mathbf{1}[Y_{1}^{*} \ge 0] \qquad Y_{2} = \mathbf{1}[Y_{2}^{*} \ge 0] \qquad (\varepsilon_{1}, \varepsilon_{2}) \perp Z$$

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• An admissible structure

$$S_0 \equiv \{h_0, F_{UX|Z}^0\} \Rightarrow F_{YX|Z}^0$$
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• These characterise the identified set.

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For all x, $P[h(X, U) \le h(X, 0.25)|x, z] \ge P[U \le 0.25|x, z]$



Averaging over X: $P[Y \le h(X, 0.25)|z] \ge 0.25$

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- (A): If $h_* \in \mathcal{H}_0$ then for all $\tau \in (0, 1)$ and $z \in \Omega$:

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- Proof:
 - If h_* is in an admissible structure delivering $F^*_{YX|Z}$ then for all $\tau\in(0,1)$ and $z\in\Omega$

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• If S_* and S_0 are observationally equivalent $F^*_{YX|Z} = F^0_{YX|Z}$.

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• **Proof**: by contradiction.

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- (C). Sharpness. If for all $\tau \in (0, 1)$ and $z \in \Omega$:

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then there exists a distribution function $F^*_{UX|Z}$ such that $S_* \equiv \{h_*, F^*_{UX|Z}\}$ is admissible and generates $F^*_{YX|Z} = F^0_{YX|Z}$ for all $z \in \Omega$.

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• **Proof**: constructive - see Annex of the paper.

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• Binary Y delivered by:

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• Notation:

$$\theta_1 \equiv p(x_1), \dots, \theta_K \equiv p(x_K)$$

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What is the set defined by

$$\left\{ h: \left(\begin{array}{c} P[Y \le h(X,\tau) | Z = z] \ge \tau \\ \\ P[Y < h(X,\tau) | Z = z] < \tau \end{array} \right) \qquad \forall \quad \tau \in (0,1), \quad z \in \Omega \right\}$$

in this case?

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• The (proposed) order of $\theta_1, \ldots, \theta_K$ is important. There are K! orderings. Suppose

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$$\{Y < h(X, \tau)\}$$
 is equal to $\{(Y = 0) \cap (p(X) < \tau)\}$

and so:

$$P[Y < h(X, \tau) | Z = z] = P[(Y = 0) \cap (p(X) < \tau) | Z = z]$$

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• For j such that $\theta_j < \tau \le \theta_{j+1}$, the event $\{p(X) < \tau\}$ occurs iff $X \in \{x_1,\ldots,x_j\}$ so

$$P[Y < h(X, \tau) | Z = z] = \sum_{k=1}^{j} \alpha_k(z) \beta_k(z)$$

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• This is less than au for all $au \in (heta_j, heta_{j+1}]$ only if

$$\sum_{k=1}^{j} \alpha_k(z) \beta_k(z) \leq \theta_j.$$

• A similar argument for the event $\{Y \le h(X, \tau)\}$ delivers

$$P[Y \le h(X,\tau)|Z=z] = \sum_{k=1}^{j} \beta_k(z) + \sum_{k=j+1}^{K} \alpha_k(z)\beta_k(z) \ge \tau$$

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• Combining, for $j \in \{1, \dots, K\}$

$$\sum_{k=1}^{j} \alpha_k(z) \beta_k(z) \leq \theta_j \leq \sum_{k=1}^{j-1} \beta_k(z) + \sum_{k=j}^{K} \alpha_k(z) \beta_k(z)$$

• These must hold for all $z \in \Omega$, so for $j \in \{1, \dots, K\}$

$$\max_{z \in \Omega} \left(\sum_{k=1}^{j} \alpha_k(z) \beta_k(z) \right) \leq \theta_j \leq \min_{z \in \Omega} \left(\sum_{k=1}^{j-1} \beta_k(z) + \sum_{k=j}^{K} \alpha_k(z) \beta_k(z) \right)$$

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- With K = 2 for each $z \in \Omega$, (but drop z from notation)

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• Swapping indexes.

$$\alpha_2\beta_2 \leq \theta_2 \leq \alpha_2\beta_2 + \alpha_1\beta_1$$

 $\alpha_2\beta_2 + \alpha_1\beta_1 \quad \leq \theta_1 \leq \quad \beta_2 + \alpha_1\beta_1$

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Calculations for a binary Y binary X example

• Here is a process for $Y \in \{0, 1\}$ and $X \in \{0, 1\}$

$$\begin{array}{rcl} Y^* &=& \alpha_0 + \alpha_1 X + \varepsilon \\ X^* &=& \beta_0 + \beta_1 Z + \eta \end{array} \\ Y = \mathbf{1}[Y^* > 0] & X = \mathbf{1}[X^* > 0] & \left[\begin{array}{c} \varepsilon \\ \eta \end{array} \right] |Z \sim N\left(\left[\begin{array}{c} 0 \\ 0 \end{array} \right], \left[\begin{array}{c} \mathbf{1} & \rho \\ \rho & \mathbf{1} \end{array} \right] \right) \end{array} \end{array}$$

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Here is an IV model

$$Y = \begin{cases} 0 & , & 0 < U \leq p(X) \\ 1 & , & p(X) < U \leq 1 \end{cases} \qquad Z \perp U \sim Unif(0,1)$$

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Consider identification of p(0), and p(1).

• The IV model is correctly specified:

$$Z \perp U \equiv \Phi(\varepsilon) \sim Unif(0,1)$$

$$Y = \begin{cases} 0 & , & 0 < U \leq \Phi(-\alpha_0 - \alpha_1 X) \\ 1 & , & \Phi(-\alpha_0 - \alpha_1 X) < U \leq 1 \end{cases}$$

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• Consider the case with ho=-0.25 and

 $\alpha_0 = 0 \qquad \alpha_1 = 0.5$ $\beta_0 = 0 \qquad \beta_1 = 1$

for which

$$p(0) = \Phi(-\alpha_0) = 0.5$$

$$p(1) = \Phi(-\alpha_0 - \alpha_1) = 0.308$$

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A parametric example: an ordered probit IV model

• Known thresholds c_1, \ldots, c_{M-1} and independence: $Z \perp U \sim Unif(0, 1)$

$$Y = \begin{cases} 1 & , & 0 < U \leq \Phi(c_1 - \alpha_0 - \alpha_1 X) \\ 2 & , & \Phi(c_1 - \alpha_0 - \alpha_1 X) < U \leq \Phi(c_2 - \alpha_0 - \alpha_1 X) \\ \vdots & & \vdots & \vdots \\ M & & \Phi(c_{M-1} - \alpha_0 - \alpha_1 X) < U \leq 1 \end{cases}$$

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• Consider the set of values of a_0 and a_1 identified by this model when:

$$Y^* = a_0 + a_1 X + \varepsilon \qquad X = b_0 + b_1 Z + \eta$$
$$Y = m, \qquad c_{m-1} < Y^* \le c_m \qquad \left[\begin{array}{c} \varepsilon \\ \eta \end{array}\right] |Z \sim N\left(\left[\begin{array}{c} 0 \\ 0 \end{array}\right], \left[\begin{array}{c} 1 & s_{\varepsilon\eta} \\ s_{\varepsilon\eta} & s_{\eta\eta} \end{array}\right]\right)$$
with: $\Omega = [-1, 1], a_0 = 0, a_1 = 1, b_0 = 0, s_{\varepsilon\eta} = 0.6, s_{\eta\eta} = 1.$

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with: $\Omega = [-1, 1], a_0 = 0, a_1 = 1, b_0 = 0, s_{\varepsilon\eta} = 0.6, s_{\eta\eta} = 1.$ We:

vary discreteness: $M \in \{5, 11, 21\}$ vary strength/support of instrument: $b_1 \in \{1, 2\}$

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M classes : $E[X|Z=z] = b_1z$



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• Intersection bounds: for each distribution $F_{YX|Z}^0$ the identified set of structural functions \mathcal{H}_0 is all h such that

$$\min_{z \in \Omega} P_0[Y \le h(X, \tau) | Z = z] \ge \tau$$

$$\max_{z \in \Omega} P_0[Y < h(X, \tau) | Z = z] < \tau$$
 for all $\tau \in [0, 1]$

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• Enumerate the set defined using an estimate, $\hat{F}^0_{YX|Z}$ - Chernozhukov, Lee & Rosen (2008).

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- Enumerate the set defined using an estimate, $\hat{F}^0_{YX|Z}$ Chernozhukov, Lee & Rosen (2008).
- Moment inequalities: for any w(z) > 0 and all $\tau \in (0, 1)$

$$E_{YXZ}[(1[Y \le h(X, \tau)] - \tau) \times w(Z)] \ge 0$$

$$E_{YXZ}[(\mathbf{1}[Y < \mathbf{h}(X, \tau)] - \tau) \times w(Z)] < 0$$

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• Andrews, Berry, Jia (2004), Rosen (2006), Pakes, Porter, Ho, Ishii (2006).

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Multivariate discrete outcomes

• $Y = (Y_1, \ldots, Y_T)$ with

 $Y_t = h_t(X, U_t)$

each h_t non-decreasing in $U_t \sim Unif(0, 1)$ and $U \equiv (U_1, \dots, U_T) \perp Z$.

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• Consider
$$S_0 \equiv \{h_1^0, \dots, h_T^*, F_{UX|Z}^0\}$$
 with copula $F_{U|Z}^0 = F_U^0$ delivering $F_{YX|Z}^0$.

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- Consider $S_0 \equiv \{h_1^0, \dots, h_T^*, F_{UX|Z}^0\}$ with copula $F_{U|Z}^0 = F_U^0$ delivering $F_{YX|Z}^0$.
- Identified set: consists of admissible

 $\{h_1^*, \ldots, h_T^*, F_U^*\}$

such that for all $au \in (0,1)^{\mathcal{T}}$, $z \in \Omega$

$$P_{0}[\bigcap_{t=1}^{T}(Y_{t} \leq h_{t}^{*}(X, \tau_{t}))|Z = z] \geq F_{U}^{*}(\tau)$$

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Binary Y, measurement error

• Impose monotone index restriction, $b(\cdot)$ is increasing

$$Y = h(\tilde{X}, U) \equiv \begin{cases} 0 & , & 0 \le U \le b(\tilde{X}'\beta) \\ 1 & , & b(\tilde{X}'\beta) < U \le 1 \end{cases} \qquad X = \tilde{X} + W$$
$$(U, W) \perp Z$$

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$$(U, W) \perp Z$$

• implies:

$$\mathbf{Y} = \left\{ \begin{array}{ll} \mathbf{0} & , & -\infty \leq & b^{-1}(U) + W'\beta & \leq X'\beta \\ \mathbf{1} & , & X'\beta < & b^{-1}(U) + W'\beta & \leq \infty \end{array} \right.$$

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Define

$$V \equiv C(b^{-1}(U) + W'\beta) \sim Unif(0,1) \perp Z$$

then

$$Y = \begin{cases} 0 & , & 0 \leq V & \leq C(X'\beta) \\ 1 & , & C(X'\beta) < V & \leq 0 \end{cases} \qquad Z \perp V \sim Unif(0,1)$$

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- The extent of the identified set depends on strength and support of instruments and the discreteness of the outcome.

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- What to do?
- Challenges include:
 - results for other models admitting multiple sources of heterogeneity, e.g. MNL type models.
 - identification catalogues: identified sets for $S = \{h, F_{UX|Z}\}$ and from this for functionals $\theta(S)$.

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