# Instrumental Variable Models for Discrete Outcomes 

Department Seminar: UIUC Economics Department

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CeMMAP \& UCL

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## Single equation IV model for discrete data

- Discrete $Y$ is determined by vector $X$ and scalar unobserved continuously distributed $U$ :

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Y=h(X, U)
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$h$ weakly monotonic in $U$, non-decreasing.

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- Two examples:
- binary $Y$, discrete $X$
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- Extensions/applications.


## Threshold crossing representation

- $Y \in\{0,1, \ldots, M\}$ determined by $X$ and $U \sim \operatorname{Unif}(0,1)$ :

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- Threshold crossing representation. Consider some $h_{0}$.

$$
h_{0}(x, u)=\left\{\begin{array}{ccrl}
0 & , & 0< & u \leq p_{0}^{0}(x) \\
1 & , & p_{0}^{0}(x)< & u \leq p_{1}^{0}(x) \\
\vdots & \vdots & \vdots & \vdots \\
M & , & p_{M-1}^{0}(x)< & u \leq 1
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- Consider a structure $S_{0} \equiv\left\{h_{0}, F_{U X \mid Z}^{0}\right\}$ with

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F_{U X \mid Z}^{0}(u, x \mid z) \equiv \operatorname{Pr}[U \leq u \cap X \leq x \mid Z=z]
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- It determines a distribution function of $Y$ and $X$ given $Z$

$$
F_{Y X \mid Z}^{0}(m, x \mid z)=F_{U X \mid Z}^{0}\left(p_{m}^{0}(x), x \mid z\right)
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- $U \Perp Z$ limits adjustment of the $U$ and $X$ arguments of admissible $F_{U X \mid Z}$ because for all $\tau, z$

$$
F_{U X \mid Z}(\tau, \infty \mid z) \equiv F_{U \mid Z}(\tau \mid z)=F_{U}(\tau)=\tau
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## Some related results:

- Continuous outcomes: Chernozhukov and Hansen (2005) and related papers.

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- Triangular models: structural equation for (continuous) $X$ :

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\begin{aligned}
& Y=h(X, U) \quad(U, V) \Perp Z \\
& X=g(X, V) \quad(U, Z
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Chesher (2003, 2005), Imbens \&Newey (2003).

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- Simultaneous models: "single equation" analysis of Tamer's (2003) entry game.

$$
\begin{aligned}
& Y_{1}^{*}=\alpha_{1} Y_{2}+Z \beta_{1}+\varepsilon_{1} \quad Y_{2}^{*}=\alpha_{2} Y_{1}+Z \beta_{2}+\varepsilon_{2} \\
& Y_{1}=1\left[Y_{1}^{*} \geq 0\right] \quad Y_{2}=1\left[Y_{2}^{*} \geq 0\right] \quad\left(\varepsilon_{1}, \varepsilon_{2}\right) \Perp Z
\end{aligned}
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## The single equation IV model: inequalities

- $Y$ is determined by observable $X$ and scalar unobservable $U$.

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- An admissible structure

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- These characterise the identified set.

For all $x, P[h(X, U) \leq h(X, 0.25) \mid x, z] \geq P[U \leq 0.25 \mid x, z]$


Averaging over $X: P[Y \leq h(X, 0.25) \mid z] \geq 0.25$

## Results concerning the identified set

- $\mathcal{H}_{0}$ is the set of structural functions, $h$, in admissible structures observationally equivalent to $S_{0} \equiv\left\{h_{0}, F_{U X \mid Z}^{0}\right\}$ :
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- (A): If $h_{*} \in \mathcal{H}_{0}$ then for all $\tau \in(0,1)$ and $z \in \Omega$ :

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- If $h_{*}$ is in an admissible structure delivering $F_{Y X \mid Z}^{*}$ then for all $\tau \in(0,1)$ and $z \in \Omega$

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- If $S_{*}$ and $S_{0}$ are observationally equivalent $F_{Y X \mid Z}^{*}=F_{Y X \mid Z}^{0}$.


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- Proof: by contradiction.


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- Let $P_{0}$ indicate probabilities taken under $F_{Y X \mid Z}^{0}$.
- (C). Sharpness. If for all $\tau \in(0,1)$ and $z \in \Omega$ :

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then there exists a distribution function $F_{U X \mid Z}^{*}$ such that $S_{*} \equiv\left\{h_{*}, F_{U X \mid Z}^{*}\right\}$ is admissible and generates $F_{Y X \mid Z}^{*}=F_{Y X \mid Z}^{0}$ for all $z \in \Omega$.

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- Proof: constructive - see Annex of the paper.


## Binary Y and discrete X

- Binary $Y$ delivered by:

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- Notation:

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\theta_{1} \equiv p\left(x_{1}\right), \ldots, \theta_{K} \equiv p\left(x_{K}\right)
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- For $k \in\{1, \ldots, K\}$ data are informative about:

$$
\alpha_{k}(z) \equiv P\left[Y=0 \mid X=x_{k}, Z=z\right] \quad \beta_{k}(z) \equiv P\left[X=x_{k} \mid Z=z\right]
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- What is the set defined by

$$
\left\{h:\binom{P[Y \leq h(X, \tau) \mid Z=z] \geq \tau}{P[Y<h(X, \tau) \mid Z=z]<\tau} \quad \forall \tau \in(0,1), \quad z \in \Omega\right\}
$$

in this case?

## The identified set

- The (proposed) order of $\theta_{1}, \ldots, \theta_{K}$ is important. There are $K$ ! orderings. Suppose

$$
0 \equiv \theta_{0}<\theta_{1} \leq \theta_{2} \leq \cdots \leq \theta_{K}<\theta_{K+1} \equiv 1
$$

## The identified set

- The (proposed) order of $\theta_{1}, \ldots, \theta_{K}$ is important. There are $K$ ! orderings. Suppose

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0 \equiv \theta_{0}<\theta_{1} \leq \theta_{2} \leq \cdots \leq \theta_{K}<\theta_{K+1} \equiv 1
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- The event

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\{Y<h(X, \tau)\} \text { is equal to }\{(Y=0) \cap(p(X)<\tau)\}
$$

and so:

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$$

- This is less than $\tau$ for all $\tau \in\left(\theta_{j}, \theta_{j+1}\right]$ only if

$$
\sum_{k=1}^{j} \alpha_{k}(z) \beta_{k}(z) \leq \theta_{j}
$$

## The identified set

- A similar argument for the event $\{Y \leq h(X, \tau)\}$ delivers

$$
P[Y \leq h(X, \tau) \mid Z=z]=\sum_{k=1}^{j} \beta_{k}(z)+\sum_{k=j+1}^{K} \alpha_{k}(z) \beta_{k}(z) \geq \tau
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$$

- Combining, for $j \in\{1, \ldots, K\}$

$$
\sum_{k=1}^{j} \alpha_{k}(z) \beta_{k}(z) \leq \theta_{j} \leq \sum_{k=1}^{j-1} \beta_{k}(z)+\sum_{k=j}^{K} \alpha_{k}(z) \beta_{k}(z)
$$

## The identified set

- These must hold for all $z \in \Omega$, so for $j \in\{1, \ldots, K\}$

$$
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$$
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\alpha_{1} \beta_{1} & \leq \theta_{1} \leq \alpha_{1} \beta_{1}+\alpha_{2} \beta_{2} \\
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$$

- Swapping indexes.

$$
\begin{aligned}
\alpha_{2} \beta_{2} & \leq \theta_{2} \leq \alpha_{2} \beta_{2}+\alpha_{1} \beta_{1} \\
\alpha_{2} \beta_{2}+\alpha_{1} \beta_{1} & \leq \theta_{1} \leq \beta_{2}+\alpha_{1} \beta_{1}
\end{aligned}
$$

## Calculations for a binary Y binary X example

- Here is a process for $Y \in\{0,1\}$ and $X \in\{0,1\}$

$$
\begin{aligned}
Y^{*} & =\alpha_{0}+\alpha_{1} X+\varepsilon \\
X^{*} & =\beta_{0}+\beta_{1} Z+\eta
\end{aligned}
$$

$$
Y=1\left[Y^{*}>0\right] \quad X=1\left[X^{*}>0\right] \quad\left[\begin{array}{l}
\varepsilon \\
\eta
\end{array}\right] \left\lvert\, Z \sim N\left(\left[\begin{array}{l}
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Y=\left\{\begin{array}{rrr}
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Consider identification of $p(0)$, and $p(1)$.

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Consider identification of $p(0)$, and $p(1)$.

- The IV model is correctly specified:

$$
\begin{gathered}
Z \Perp U \equiv \Phi(\varepsilon) \sim \operatorname{Unif}(0,1) \\
Y=\left\{\begin{aligned}
0, & 0<U \leq \Phi\left(-\alpha_{0}-\alpha_{1} X\right) \\
1, & \Phi\left(-\alpha_{0}-\alpha_{1} X\right)
\end{aligned}\right) \quad U \leq 1
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\rho & 1
\end{array}\right]\right)\right.
\end{aligned}
$$

- Consider the case with $\rho=-0.25$ and

$$
\begin{array}{cc}
\alpha_{0}=0 & \alpha_{1}=0.5 \\
\beta_{0}=0 & \beta_{1}=1
\end{array}
$$

for which

$$
\begin{aligned}
p(0) & =\Phi\left(-\alpha_{0}\right)=0.5 \\
p(1) & =\Phi\left(-\alpha_{0}-\alpha_{1}\right)=0.308
\end{aligned}
$$















## A parametric example: an ordered probit IV model

- Known thresholds $c_{1}, \ldots, c_{M-1}$ and independence: $Z \Perp U \sim \operatorname{Unif}(0,1)$

$$
Y=\left\{\begin{array}{rrrl}
1 & , & 0< & U \\
2 & , & \Phi\left(c_{1}-\alpha_{0}-\alpha_{1} X\right)<\Phi\left(c_{1}-\alpha_{0}-\alpha_{1} X\right) \\
\vdots & \vdots & \vdots \Phi\left(c_{2}-\alpha_{0}-\alpha_{1} X\right) \\
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& \left.-\alpha_{1} X\right)< & U & \leq 1
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$$

- Consider the set of values of $a_{0}$ and $a_{1}$ identified by this model when:

$$
\begin{aligned}
& Y^{*}=a_{0}+a_{1} X+\varepsilon \quad X=b_{0}+b_{1} Z+\eta \\
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& \text { with: } \Omega=[-1,1], a_{0}=0, a_{1}=1, b_{0}=0, s_{\varepsilon \eta}=0.6, s_{\eta \eta}=1 .
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with: $\Omega=[-1,1], a_{0}=0, a_{1}=1, b_{0}=0, s_{\varepsilon \eta}=0.6, s_{\eta \eta}=1$.

- We:

$$
\text { vary discreteness: } M \in\{5,11,21\}
$$

vary strength/support of instrument: $b_{1} \in\{1,2\}$

Mclasses: $E[X \mid Z=z]=b_{1} z$


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## Estimation

- Intersection bounds: for each distribution $F_{Y X \mid Z}^{0}$ the identified set of structural functions $\mathcal{H}_{0}$ is all $h$ such that

$$
\left.\begin{array}{l}
\min _{z \in \Omega} P_{0}[Y \leq h(X, \tau) \mid Z=z] \geq \tau \\
\max _{z \in \Omega} P_{0}[Y<h(X, \tau) \mid Z=z]<\tau
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- Moment inequalities: for any $w(z)>0$ and all $\tau \in(0,1)$

$$
\begin{gathered}
E_{Y X Z}[(1[Y \leq h(X, \tau)]-\tau) \times w(Z)] \geq 0 \\
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$$

- Andrews, Berry, Jia (2004), Rosen (2006), Pakes, Porter, Ho, Ishii (2006).


## Multivariate discrete outcomes

- $Y=\left(Y_{1}, \ldots, Y_{T}\right)$ with

$$
Y_{t}=h_{t}\left(X, U_{t}\right)
$$

each $h_{t}$ non-decreasing in $U_{t} \sim \operatorname{Unif}(0,1)$ and $U \equiv\left(U_{1}, \ldots, U_{T}\right) \Perp Z$.

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- Consider $S_{0} \equiv\left\{h_{1}^{0}, \ldots h_{T}^{*}, F_{U X \mid Z}^{0}\right\}$ with copula $F_{U \mid Z}^{0}=F_{U}^{0}$ delivering $F_{Y X \mid Z}^{0}$.


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- Identified set: consists of admissible

$$
\left\{h_{1}^{*}, \ldots, h_{T}^{*}, F_{U}^{*}\right\}
$$

such that for all $\tau \in(0,1)^{T}, z \in \Omega$

$$
P_{0}\left[\bigcap_{t=1}^{T}\left(Y_{t} \leq h_{(<)}^{*}\left(X, \tau_{t}\right)\right) \mid Z=z\right] \underset{(<)}{\geq} F_{U}^{*}(\tau)
$$

## Binary Y , measurement error

- Impose monotone index restriction, $b(\cdot)$ is increasing

$$
\begin{gathered}
Y=h(\tilde{X}, U) \equiv\left\{\begin{array}{cc}
0, & 0 \leq U \leq b\left(\tilde{X}^{\prime} \beta\right) \quad X=\tilde{X}+W \\
1, & b\left(\tilde{X}^{\prime} \beta\right)<U \leq 1
\end{array}(U, W) \Perp z\right.
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(U, W) \Perp Z
\end{array}\right.
\end{gathered}
$$

- implies:

$$
Y=\left\{\begin{array}{lll}
0 & , & -\infty \leq b^{-1}(U)+W^{\prime} \beta \leq X^{\prime} \beta \\
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1 & , & X^{\prime} \beta<b^{-1}(U)+W^{\prime} \beta \leq \infty
\end{array}\right.
$$

- Define

$$
V \equiv C\left(b^{-1}(U)+W^{\prime} \beta\right) \sim U \operatorname{Uif}(0,1) \Perp z
$$

then

$$
Y=\left\{\begin{aligned}
0, & 0 \leq V \leq C\left(X^{\prime} \beta\right) \\
1 & , \quad C\left(X^{\prime} \beta\right)<V \leq 0
\end{aligned} \quad Z \Perp V \sim \operatorname{Unif}(0,1)\right.
$$

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- Extensions: multivariate outcomes, measurement error.
- What to do?


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- identification catalogues: identified sets for $S=\left\{h, F_{U X \mid Z}\right\}$ and from this for functionals $\theta(S)$.

