The Ignorant Monopolist Redux

Roger Koenker[†]

[†]*Department of Economics, University College London, 30 Gordon Street, London, WC1H 0AX, UK*
E-mail: `r.koenker@ucl.ac.uk`

Summary   The classical problem of the monopolist faced with an unknown demand curve is considered in a simple stochastic setting. Sequential pricing strategies designed to maximize discounted profits are shown to converge sufficiently rapidly that they leave the monopolist ignorant about all but the most local features of demand. The failure of the monopolist to "learn" his demand curve would seem to call into question some standard assumptions about agents' grasp of their economic environment.

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1. INTRODUCTION

Clower (1959) describes the dilemma of the monopolist facing an unknown, but non-stochastic, demand function. By a sequence of Newton-like steps he gradually finds his way to a profit maximizing output/pricing policy, and yet is left ignorant of the global features of his economic environment, learning only about its local appearance near the profit maximizing price. Since Clower there have been many investigations that introduced stochastic features into this monopoly predicament. The monopolist is always torn between the desire to maximize his (presumably discounted) stream of profits and his wish to learn more about his economic environment. Learning is presumably helpful in the relentless quest for better profits, but too much experimentation leads to sacrifices in valuable early profits.

Following Grossman et al. (1977) most of the literature has posed such problems in the context of a dynamic programming formulation with Bayesian updating of the monopolist's beliefs about demand and cost. See, e.g. Easley and Kiefer (1988) and Nyarko (1991). Even in very simple environments with known parametric forms for demand and cost, this approach yields computationally challenging problems. In this brief note I would like to consider a much simpler class of sequential design strategies employing the stochastic approximation methods introduced by Kiefer and Wolfowitz (1952). I will show that although these methods can achieve a certain optimality from a profitability perspective, they fail miserably in the secondary objective of quenching the monopolist's curiosity about the precise form of his demand curve.

2. THE MODEL

Suppose that our monopolist faces the log-quadratic demand function,

$$D(p, u) = \exp\{\alpha_0 + \alpha_1 \log(p) + \alpha_2 (\log(p))^2 + u\}.$$

For added simplicity we will take the random demand shocks, u, to be iid Gaussian with mean zero, and constant variance, σ^2. Cost will be assumed to be non-stochastic and again log-quadratic,

$$C(x) = \exp\{\beta_0 + \beta_1 \log(x) + \beta_2 (\log(x))^2\},$$

The monopolist's objective is to choose a sequence of prices, $\{p_t, t = 1, 2, \cdots\}$, to maximize discounted profit,

$$R(p) = \sum_{t=1}^{\infty} \delta^t \pi(p_t) \equiv \sum_{t=1}^{\infty} \delta^t (p_t X_t - C(X_t)),$$

where $X_t = D(p_t, u_t)$. Since cost is nonstochastic it is easy to learn its parameters, so we will treat them as known from the outset and focus on the unknown demand parameters and the choice of a profit maximizing pricing sequence.

3. STOCHASTIC APPROXIMATION

Robbins and Monro (1951) introduced a recursive strategy for finding a zero crossing of the expected value of a random function; Kiefer and Wolfowitz (1952) noted that similar methods could be employed to find the argmax of an unknown regression function. The recursion may be viewed as an adaptive variant of the classical Newton method. Given an initial guess p_0 of $p^* = \text{argmax } \mathbb{E}\pi(p)$ we let

$$p_{t+1} = p_t + a_t(\pi(p_t + c_t) - \pi(p_t - c_t))/(2c_t).$$

The sequences $\{a_t\}$ and $\{c_t\}$, as shown by Fabian (1971), are optimally taken to be $a_t = a_0 t^{-1}$ and $c_t = c_0 t^{-1/6}$ respectively. Assuming that expected profit,

$$\pi_0(p) = \mathbb{E}\pi(p)$$

is sufficiently smooth in a neighborhood of a unique maximizing value, p^*, Fabian establishes that, with his choices,

$$n^{1/3}(p_n - p^*) \rightsquigarrow \mathcal{N}(\mu(p^*), \sigma^2(p^*)),$$

where

$$\mu(p) = \frac{-a_0 c_0^2 \pi_0'''(p)}{3(2\pi_0''(p) - 2/3)}$$

and

$$\sigma^2(p) = \frac{a_0^2 \sigma^2(p)}{2c_0^2(2a_0\pi_0''(p) - 2/3)}.$$

Optimal choices of the constants, a_0 and c_0 are shown by Abdelhamid (1973) to be, $a_0 = (\pi_0''(p^*))^{-1}$ and $c_0 = (3\sigma^2(p^*))/\pi_0'''(p^*))^{1/6}$, where $\sigma^2(p)$ denotes the conditional variance of profits at p^*.

The cube root convergence of the pricing sequence is obviously unfortunate from a profit maximizing perspective; the stereotypic greedy monopolist may feel entitled to the parametric rate $O(1/\sqrt{n})$. Fabian's results reveal that even after adopting the optimal choices of the $\{a_t\}$ and $\{c_t\}$ sequences, we obtain a biased limiting distribution, after normalization, for the $\{p_t\}$ sequence. This is an inevitable consequence of the non-parametric nature of problem of finding a root of $\pi_0'(p) = 0$; in the original Robbins-Monro setting one has $\sqrt{n}$ convergence to a normal limit without the possible asymptotic bias effect induced when $\pi_0'''(p^*) \neq 0$. For our monopolist both bias and variance of the pricing sequence decline like $n^{-1/3}$ as p^* is approached. Stochastic approximation must, however, be gauged a huge success for profit maximization compared to its ability to reveal the precise form of the demand curve.

4. LEARNING ABOUT DEMAND

Even the limited success we can achieve in finding p^* poses some serious problems for the curious monopolist who would like to learn more about his demand function. After n steps of the stochastic approximation iteration, our monopolist has $3n$ observations on the demand curve, but the experimental design of these observations makes them less than ideal for learning about demand.

Before considering the quadratic specification, suppose that the demand were known to be linear in price. In this case the precision of the slope parameter in a demand regression is proportional to the sum of squares of the centered design observations. Let $z_t = \log(p_t/p^*)$, since z_t is converging to zero at the rate, $n^{-1/3}$, we have

$$\sum_{t=1}^{n} z_t^2 \approx K \sum_{t=1}^{n} t^{-2/3} \approx K n^{1/3},$$

and consequently our estimated slope parameter of the demand function converges at the phlegmatic rate of $O_p(n^{-1/6})$.

The quadratic case is even more troubling since the quadratic coefficient of the demand curve is estimated with precision proportional to the fourth moment of $\log p$, and thus,

$$\sum_{t=1}^{n} z_t^4 \approx K \sum_{t=1}^{n} t^{-4/3} \approx 3.60K < \infty,$$

and consequently the quadratic coefficient *cannot* be consistently estimated, as explained in further detail in the Appendix. Despite the conscientious efforts of the monopolist to profit maximize, or indeed because of these efforts, he is doomed to ignorance about his own demand environment. For those imagining a more ambitious nonparametric demand environment the situation is even more bleak.

4.1. A Numerical Example

To illustrate the foregoing situation we consider the following explicit version of the model:

$$D(p, u) = \exp\{5 - 0.5\log(p) - 0.35(\log(p))^2 + u\},$$

where u is iid Gaussian with mean zero, and constant variance, $\sigma^2 = 0.0025$. Cost is

$$C(x) = \exp\{2 + 0.3\log(x) + 0.1(\log(x))^2\},$$

In Figure 1 we see a scatter plot of 1000 realizations of profit evaluated at uniformly distributed p_t over the interval $[3, 60]$. Superimposed is an estimate of the expected profit function, $\mu_0(p)$. Of course, any self-respecting monopolist would be crazy to spread his pricing strategy so thinly over such a wide range of prices far from the optimum, but we not are so constrained.

In Figure 2 we illustrate a more realistic sequence of prices and their corresponding demand realizations; these prices arise from 100 iterations of the basic Kiefer-Wolfowitz procedure. As described above, every third p_t is a "real" iterate and the intermediate evaluations of demand are used to compute the estimated derivative. In the figure we have connected the former points to illustrate the Kiefer-Wolfowitz trajectory. Clearly there is a rapid ascent to the profit maximizing value, and the isolated points on either side of this value illustrate the continued effort to refine this estimate with very gradually
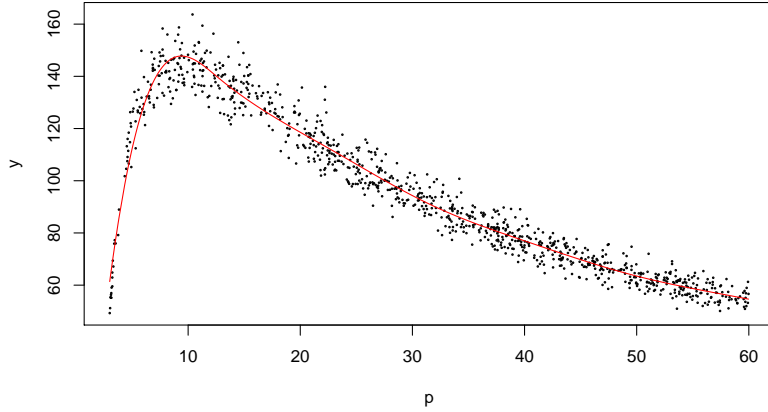
Figure 1. Realized profit for 1000 uniformly distributed p_t on the interval $[3, 60]$. An estimate of the expected profit function is superimposed on the scatterplot.

declining bandwidths for the derivative estimates. The starting value for this example was $p_0 = 15$, so we are able to see reasonably well what demand looks like for $p \in [8, 15]$, but this is not sufficient to allow us to estimate well the demand curve. Least squares on the full sample of 300 observations yields:

$$\log x_t = \underset{(0.48)}{4.58} - \underset{(0.41)}{0.16} \log p_t - \underset{(0.09)}{0.25} (\log p_t)^2.$$

Inference based on these estimates is fine, the true parameters are covered by their respective confidence intervals, but even after extensive continued iteration, because of the concentration of the design points around p^*, we fail to get convergent estimates of the demand parameters.

Results in this section can be reproduced by downloading the R package `monopoly` from `http://www.econ.uiuc.edu/~roger/research/monopoly/mono.html`. In particular, Figures 1 and 2 can be reproduced by running `demo(fig1)` and `demo(fig2)`.

5. RAMIFICATIONS

Fabian (1967) has proposed a more elaborate variant of the Kiefer-Wolfowitz procedure whose convergence rate approaches the parametric, $O_p(1/\sqrt{n})$ ideal provided that the underlying $\mu_0(p)$ function is sufficiently smooth near p^*. This can be regarded as a stochastic approximation variant of the more familiar use of higher order kernels to do bias removal for non-parametric density estimation. Ruppert (1991) notes that there is little practical (simulation) experience regarding this procedure. It may be noted that any "improved" performance of the sequence $\{p_t\}$ from the profitability perspective,
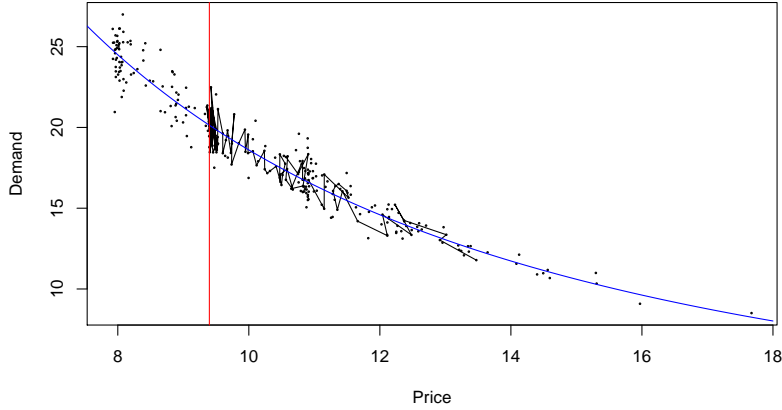
Figure 2. Realized price and demand for 100 iterations of the Kiefer-Wolfowitz process. The trajectory of the central iterates is depicted in black and the isolated points are those required to compute the difference quotients. A least squares estimate of the demand curve is over plotted in blue, and the final estimate of the profit maximizing price is depicted by the vertical red line.

accentuates the difficulties from a demand analysis perspective. In this case we have,

$$\sum_{t=1}^{n} z_t^2 = K \sum_{t=1}^{n} t^{-1} \approx K \log n,$$

and even the slope coefficient for demand is barely estimable, converging at the rate $O_p(1/\sqrt{\log n})$.

Of course knowledge of the parametric form of cost and demand, iid error for the random component of demand, stationary of the demand process and a variety of other simplifying assumptions can be easily called into question. But relaxing such conditions only leaves the ignorant monopolist deeper in the dark.

Since stochastic approximation purports to achieve optimal rates of convergence for the profit maximizing sequence of prices in problems of the foregoing type, it would be of interest to see whether Bayesian updating could improve upon the simple methods described above, and if so by how much. To the extent that improvements in the monopolist's pricing policies are made, they only further deepen his ignorance about demand.

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REFERENCES

Abdelhamid, S. (1973). Transformation of observations in stochastic approximation. *Annals of Statistics 1*, 1158–1174.

Clower, R. (1959). Some theory of an ignorant monopolist. *Economic Journal 69*, 705–716.

Easley, D. and N. Kiefer (1988). Controlling a stochastic process with unknown parameters. *Econometrica 56*, 1045–1064.

Fabian, V. (1967). Stochastic approximation of minima with improved asymptotic speed. *Annals of Mathematical Statistics 38*, 191–200.

Fabian, V. (1971). Stochastic approximation. In J. Rustagi (Ed.), *Optimizing Methods in Statistics*, pp. 439–470. Academic Press: New York.

Grossman, S., R. Kihlstrom, and L. Mirman (1977). A Bayesian approach to the production of information and learning by doing. *The Review of Economic Studies 44*, 533–547.

Kiefer, J. and J. Wolfowitz (1952). Stochastic estimation of the maximum of a regression function. *Annals of Mathematical Statistics 23*, 462–466.

Lai, T. and H. Robbins (1977). Strong consistency of least-squares estimates in regression models. *PNAS 74*, 2667–2669.

Nyarko, Y. (1991). On the convergence of Bayesian posterior processes in linear economic models. *J. of Economic Dynamics and Control 15*, 687–713.

Robbins, H. and S. Monro (1951). A stochastic approximation method. *Annals of Mathematical Statistics 22*, 400–407.

Ruppert, D. (1991). Stochastic approximation. In B. Ghosh and P. Sen (Eds.), *Handbook in Sequential Analysis*, pp. 503–529. Dekker: New York.

A. DESIGN CONDITIONS FOR CONVERGENCE OF THE LEAST SQUARES REGRESSION ESTIMATOR

Consider the bivariate linear regression model,

$$y_i = \alpha + \beta x_i + u_i \quad i = 1, \cdots, n,$$

with iid $\{u_i\}$ and $\mathbb{E}u_1 = 0$ and $\mathbb{E}u_1^2 \log(1 + |u_i|^r) < \infty$ for some $r > 1$. A necessary as well as sufficient condition for the least squares estimator $\hat{\beta}_n$ to be consistent for β is that,

$$\lim_{n \to \infty} A_n \equiv \lim_{n \to \infty} \sum (x_i - \bar{x})^2 = \infty.$$

When this condition holds, Lai and Robbins (1977) prove that for every $\delta > 0$,

$$\lim_{n \to \infty} \frac{\sqrt{A_n}(\hat{\beta} - \beta)}{(\log A_n)^{(1+\delta)/2}} = 0,$$

with probability one, so $\hat{\beta}$ converges to β essentially at rate $O_p(1/\sqrt{A_n})$. This can be strengthened to limiting normality for $\sqrt{A_n}(\hat{\beta} - \beta)$ provided that $\max |x_i|/\sqrt{A_n} \to 0$, which assures the Lindeberg condition.

For the quadratic model,

$$y_i = \alpha + \beta x_i + \gamma x_i^2 + u_i \quad i = 1, \cdots, n,$$

we can project $z = (x_i^2)$ onto the orthogonal complement of $X = (1, x_i)$, say $\tilde{z} = M_X z,$

thus reducing the question of consistency of the least squares $\hat{\gamma}$ to the requirement that $\|\tilde{z}\|^2 \to \infty$. But

$$\|\tilde{z}\|^2 \leq \|z\|^2 \equiv \sum_{i=1}^{n} x_i^4,$$

so, for example, when $x_i = x_0 + \xi_i$ with $\xi_i = O_p(i^{-1/3})$, as in the monopoly problem described above, $\|\tilde{z}\|^2$ fails to diverge, and consequently $\hat{\gamma}$ is inconsistent.