Quantile Regression Computation Outside, Inside and Proximal

Roger Koenker

University of Illinois, Urbana-Champaign

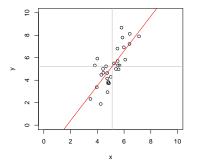
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The Origin of Regression - Regression Through the Origin

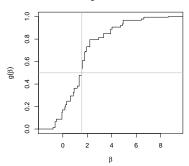
Find the line with mean residual zero that minimizes the sum of absolute residuals.



Problem: $\min_{\alpha,\beta} \sum_{i=1}^{n} |y_i - \alpha - x_i\beta|$ s.t. $\bar{y} = \alpha + \bar{x}\beta$.

Boscovich/Laplace Methode de Situation

Algorithm: Order the n candidate slopes: $b_i = (y_i - \bar{y})/(x_i - \bar{x})$ denoting them by $b_{(i)}$ with associated weights $w_{(i)}$ where $w_i = |x_i - \bar{x}|$. Find the weighted median of these slopes. Reduces the problem to (partial) sorting.

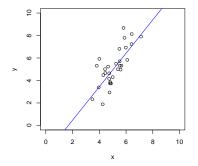


Weighted EDF Plot

Edgeworth's (1888) Plural Median

What if we want to estimate both α and β by median regression?

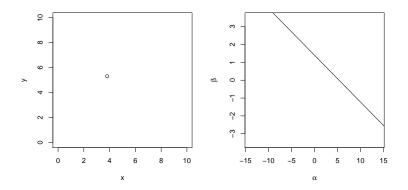
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Edgeworth's (1888) Dual Plot: Anticipating Simplex

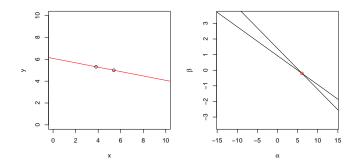
Points in sample space map to lines in parameter space.

$$(\mathbf{x}_{i}, \mathbf{y}_{i}) \mapsto \{(\alpha, \beta) : \alpha = \mathbf{y}_{i} - \mathbf{x}_{i}\beta\}$$

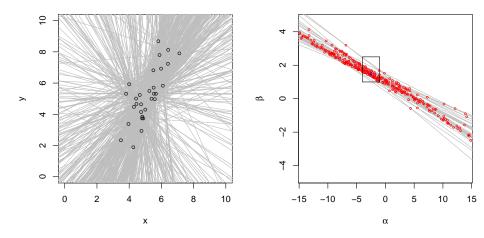


Edgeworth's (1888) Dual Plot: Anticipating Simplex

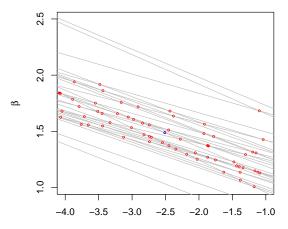
Lines through pairs of points in sample space map to points in parameter space.



Edgeworth's (1888) Dual Plot: Anticipating Simplex All pairs of observations produce $\binom{n}{2}$ points in dual plot.



Edgeworth's (1888) Dual Plot: Anticipating Simplex Follow path of steepest descent through vertices in the dual plot.



Barrodale-Roberts (1974) Implementation of Edgeworth

```
rgx<- function(x, y, tau = 0.5, max.it = 50) { # Barrodale and Roberts -- lite
        p \leftarrow ncol(x); n \leftarrow nrow(x)
        h <- sample(1:n, size = p) #Phase I -- find a random (!) initial basis
        it <- 0
        repeat {
                it <- it + 1
                Xhinv <- solve(x[h, ])</pre>
                bh <- Xhinv %*% v[h]
                rh <- v - x %*% bh
        #find direction of steepest descent along one of the edges
                g <- - t(Xhinv) %*% t(x[ - h, ]) %*% c(tau - (rh[ - h] < 0))
                g <- c(g + (1 - tau), - g + tau)
                ming <- min(g)
                if(ming >= 0 || it > max.it) break
                h.out <- seq(along = g)[g == ming]
                sigma <- ifelse(h.out <= p. 1, -1)
                if(sigma < 0) h.out <- h.out - p
                d <- sigma * Xhinv[, h.out]
        #find step length by one-dimensional wouantile minimization
                xh <- x %*% d
                step <- wquantile(xh, rh, tau)</pre>
                h.in <- step$k
                h <- c(h[ - h.out], h.in)
        3
        if(it > max.it) warning("non-optimal solution: max.it exceeded")
        return(bh)
}
```

Splitting the QR "residual" into positive and negative parts, yields the primal linear program,

 $\min_{(\mathfrak{b},\mathfrak{u},\nu)}\{\tau\mathbf{1}^{\top}\mathfrak{u}+(1-\tau)\mathbf{1}^{\top}\nu\mid X\mathfrak{b}+\mathfrak{u}-\nu-\mathfrak{y}\in\{0\},\quad (\mathfrak{b},\mathfrak{u},\nu)\in\mathsf{R}^p\times\mathsf{R}^{2n}_+\}.$

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with dual program:

$$\max_{d} \{ \boldsymbol{y}^{\top} \boldsymbol{d} \mid \boldsymbol{X}^{\top} \boldsymbol{d} \in \{\boldsymbol{0}\}, \quad \tau \boldsymbol{1} - \boldsymbol{d} \in \boldsymbol{\mathsf{R}}^{n}_{+}, \quad (\boldsymbol{1} - \tau)\boldsymbol{1} + \boldsymbol{d} \in \boldsymbol{\mathsf{R}}^{n}_{+} \},$$

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with dual program:

$$\begin{split} \max_{d} \{ y^{\top} d \mid X^{\top} d \in \{0\}, \quad \tau 1 - d \in \mathsf{R}^{n}_{+}, \quad (1 - \tau) 1 + d \in \mathsf{R}^{n}_{+} \}, \\ \max_{d} \{ y^{\top} d \mid X^{\top} d = 0, \ d \in [\tau - 1, \tau]^{n} \}, \end{split}$$

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Quantile Regression Dual

The dual problem for quantile regression may be formulated as:

$$\max_{a} \{ \boldsymbol{y}^{\top} \boldsymbol{\alpha} | \boldsymbol{X}^{\top} \boldsymbol{\alpha} = (1 - \tau) \boldsymbol{X}^{\top} \boldsymbol{1}, \ \boldsymbol{\alpha} \in [0, 1]^n \}$$

What do these $\hat{a}_i(\tau)$'s mean statistically?

They are regression rank scores (Gutenbrunner and Jurečková (1992)):

$$\hat{a}_{i}(\tau) \in \begin{cases} \{1\} & \text{if} \quad y_{i} > x_{i}^{\top} \hat{\beta}(\tau) \\ (0,1) & \text{if} \quad y_{i} = x_{i}^{\top} \hat{\beta}(\tau) \\ \{0\} & \text{if} \quad y_{i} < x_{i}^{\top} \hat{\beta}(\tau) \end{cases}$$

The integral $\int \hat{a}_i(\tau) d\tau$ is something like the rank of the ith observation. It answers the question: On what quantile does the ith observation lie? Fundamental to the construction of linear rank statistics for regression.

Linear Programming: The Inside Story

The Simplex Method (Edgeworth/Dantzig/Kantorovich) moves from vertex to vertex on the outside of the constraint set until it finds an optimum.

Interior point methods (Frisch/Karmarker/et al) take Newton type steps toward the optimal vertex from inside the constraint set.

Linear Programming: The Inside Story

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A toy problem: Given a polygon inscribed in a circle, find the point on the polygon that maximizes the sum of its coordinates:

$$\max\{e^{\top}u|A^{\top}x=u,\ e^{\top}x=1,\ x\geqslant 0\}$$

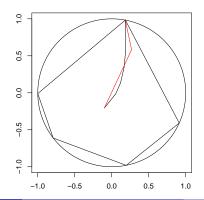
were *e* is vector of ones, and A has rows representing the n vertices. Eliminating u, setting c = Ae, we can reformulate the problem as:

$$\max\{c^{\top}x|e^{\top}x=1, \quad x \ge 0\},$$

Toy Story: From the Inside

Simplex goes around the outside of the polygon; interior point methods tunnel from the inside, solving a sequence of problems of the form:

$$\max\{c^{\top}x + \mu \sum_{i=1}^{n} \log x_i | e^{\top}x = 1\}$$

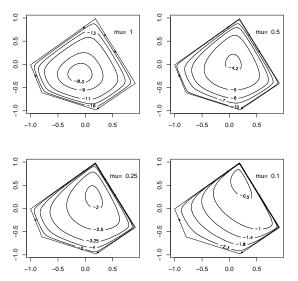


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Quantile Regression Computation

Toy Story: From the Inside

By letting $\mu \to 0$ we get a sequence of smooth problems whose solutions approach the solution of the LP:



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Mehrotra Primal-Dual Predictor-Corrector Algorithm

The algorithms implemented in my R package quantreg are based on Mehrotra's (1992) Predictor-Corrector approach. Although somewhat more complicated than prior methods it has several advantages:

- Better stability and efficiency due to better central path following,
- Easily generalized to incorporate linear inequality constraints.
- Easily generalized to exploit sparsity of the design matrix.
- Preprocessing can improve performance in large n small p problems.

These features are all incorporated into various versions of the algorithm in quantreg, and coded in Fortran.

A Model of Childhood Malnutrition in India

fit <- $rqss(cheight \sim qss(cage, lambda = lam[1]) + qss(bfed, lambda = lam[2]) + qss(mage, lambda = lam[3]) + qss(mbmi, lambda = lam[4]) + qss(sibs, lambda = lam[5]) + qss(medu, lambda = lam[6]) + qss(fedu, lambda = lam[7]) + csex + ctwin + cbirthorder + munemployed + mreligion + mresidence + deadchildren + wealth + electricity + radio + television + frig + bicycle + motorcycle + car + tau = 0.10, method = "lasso", lambda = lambda, data = india)$

- The seven coordinates of lam control the smoothness of the nonparametric components via total variation penalties,
- lambda controls the (lasso) shrinkage of the linear coefficients.
- The estimated model has roughly 40,000 "observations", including the penalty contribution, and has 2201 parameters.
- Fitting for a single choice of λ's takes approximately 5 seconds. Sparsity of the design matrix is critical to efficient Cholesky factorization at each interior point iteration.

Proximal Algorithms for Large p Problems

Given a closed, proper convex function $f: R^n \to R \cup \{\infty\}$ the proximal operator, $P_f: R^n \to R^n$ of f is defined as,

$$P_f(\nu) = \operatorname{argmin}_{\chi} \{ f(\chi) + \frac{1}{2} \| \chi - \nu \|_2^2 \}.$$

View v as an initial point and $P_f(v)$ as a half-hearted attempt to minimize f, while constrained not to venture too far away from v.

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View ν as an initial point and $P_f(\nu)$ as a half-hearted attempt to minimize f, while constrained not to venture too far away from ν .

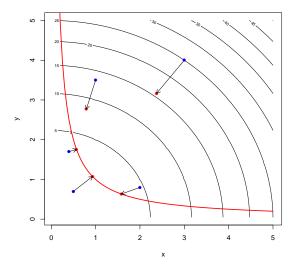
The corresponding Moreau envelope of f is

$$M_{f}(\nu) = \inf_{x} \{f(x) + \frac{1}{2} \|x - \nu\|_{2}^{2}\}.$$

thus evaluating M_f at $\nu = x$ we have,

$$M_{f}(x) = f(P_{f}(x)) + \frac{1}{2} ||x - P_{f}(x)||_{2}^{2}$$
.

A Toy Example:



Proximal Operators as (Regularized) Gradient Steps

Rescaling f by $\lambda \in \mathbb{R}$, $M_{\lambda f}(x) = f(P_{\lambda f}(x)) + \frac{1}{2\lambda} \|x - P_{\lambda f}(x)\|_2^2 \}.$ so $\nabla M_{\lambda f}(x) = \lambda^{-1} (x - P_{\lambda f}(x)),$

or

$$P_{\lambda f}(x) = x - \lambda \nabla M_{\lambda f}(x).$$

So $P_{\lambda f}$ may be interpreted as a gradient step of length λ for $M_{\lambda f}$.

Proximal Operators as (Regularized) Gradient Steps

Rescaling f by $\lambda \in R$,

$$\mathsf{M}_{\lambda \mathsf{f}}(\mathsf{x}) = \mathsf{f}(\mathsf{P}_{\lambda \mathsf{f}}(\mathsf{x})) + \frac{1}{2\lambda} \|\mathsf{x} - \mathsf{P}_{\lambda \mathsf{f}}(\mathsf{x})\|_2^2 \}.$$

so

$$\nabla M_{\lambda f}(\mathbf{x}) = \lambda^{-1}(\mathbf{x} - P_{\lambda f}(\mathbf{x})),$$

or

$$P_{\lambda f}(x) = x - \lambda \nabla M_{\lambda f}(x).$$

So $P_{\lambda f}$ may be interpreted as a gradient step of length λ for $M_{\lambda f}$. Unlike f, which may have a nasty subgradient, M_f has a nice gradient:

$$M_{f} = (f^{*} + \frac{1}{2} \| \cdot \|_{2}^{2})^{*}$$

where $f^*(y) = \sup_x \{y^\top x - f(x)\}$ is the convex conjugate of f.

Proximal Operators and Fixed Point Iteration

The gradient step interpretation of P_f suggests the fixed point iteration:

$$\mathbf{x}^{k+1} = \mathsf{P}_{\lambda \mathsf{f}}(\mathbf{x}^k).$$

While this may not be a contraction, it is "firmly non-expansive" and therefore convergent.

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In additively separable problems of the form

$$\min_{\mathbf{x}} \{ f(\mathbf{x}) + g(\mathbf{x}) \},\$$

with f and g convex, this may be extended to the ADMM algorithm:

$$\begin{split} \mathbf{x}^{k+1} &= \mathsf{P}_{\lambda \mathsf{f}}(z^k - \mathbf{u}^k) \\ z^{k+1} &= \mathsf{P}_{\lambda \mathsf{g}}(\mathbf{x}^k - \mathbf{u}^k) \\ \mathbf{u}^{k+1} &= (\mathbf{u}^k + \mathbf{x}^k - z^k) \end{split}$$

Alternating Direction Method of Multipliers, Parikh and Boyd (2013).

The Proximal Operator Graph Solver

A further extension that encompasses many currently relevant statistical problems is:

$$\min_{(\mathbf{x},\mathbf{y})} \{ \mathbf{f}(\mathbf{y}) + \mathbf{g}(\mathbf{x}) \mid \mathbf{y} = \mathbf{A}\mathbf{x} \},\$$

where (x, y) is constrained to the graph $\mathcal{G} = \{(x, y) \in \mathbb{R}^{n+m} \mid y = Ax\}$. The modified ADMM algorithm becomes:

$$\begin{split} (x^{k+1/2}, y^{k+1/2}) &= (\mathsf{P}_{\lambda g}(x^k - \tilde{x}^k), \mathsf{P}_{\lambda f}(y^k - \tilde{y}^k)) \\ (x^{k+1}, y^{k+1}) &= \Pi_A(x^{k+1/2} - \tilde{x}^k, y^{k+1/2} - \tilde{y}^k) \\ (\tilde{x}^{k+1}, \tilde{y}^{k+1}) &= (\tilde{x}^k + x^{k+1/2} - x^{k+1}, \tilde{y}^{k+1/2} + y^{k+1/2} - y^{k+1}) \end{split}$$

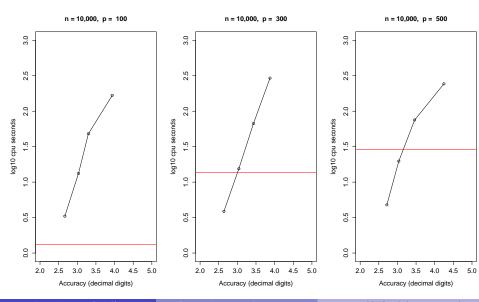
where Π_A denotes the (Euclidean) projection into graph \mathcal{G} . This has been elegantly implemented by Fougner and Boyd (2015) and made available by Fougner in the R package POGS.

When Is POGS Most Attractive?

• f and g must:

- Be closed, proper convex
- Be additively (block) separable
- Have easily computable proximal operators
- A should be:
 - Not too thin
 - Not too sparse
- Other Problem Aspects
 - Available parallelizable hardware, cluster, GPUs, etc.
 - Not too stringent accuracy requirement

POGS Performance – Large p Quantile Regression



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Quantile Regression Computation

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Global Quantile Regression?

Usually quantile regression is local, so solutions,

$$\hat{\boldsymbol{\beta}}(\tau) = \mathsf{argmin}_{\boldsymbol{b} \in \mathsf{R}^p} \sum_{i=1}^n \rho_\tau(\boldsymbol{y}_i - \boldsymbol{x}_i^\top \boldsymbol{b})$$

are sensitive only to $\{y_i\}$ near $Q(\tau|x_i),$ the τth conditional quantile function of $Y_i|X=x_i.$

But recently there has been more interest in jointly estimating several $\beta(\tau_i)$:

$$\{\hat{\boldsymbol{\beta}}(\tau) \mid \tau \in \boldsymbol{\mathfrak{T}}\} = \text{argmin} \sum_{\tau \in \boldsymbol{\mathfrak{T}}} \sum_{i=1}^n w_\tau \rho_\tau(\boldsymbol{y}_i - \boldsymbol{x}_i^\top \boldsymbol{b}_\tau)$$

This is sometimes called "composite quantile regression" as in Zou and Yuan (2008). Constraints need to be imposed on the $\beta(\tau)$ otherwise the problem separates.

Example 1: Choquet Portfolios

Bassett, Koenker and Kordas (2004) proposed estimating portfolio weights $\pi \in \mathsf{R}^p$ by solving:

$$\min_{\pi \in \mathsf{R}^{\mathsf{p}}, \ \xi \in \mathsf{R}^{\mathsf{m}}} \{ \sum_{k=1}^{\mathfrak{m}} \sum_{i=1}^{\mathfrak{n}} w_{\tau_k} \rho_{\tau_k} (x_i^\top \pi - \xi_{\tau_k}) \mid \bar{x}^\top \pi = \mu_0 \}$$

where $x_i \in \mathsf{R}^p : i = 1, \cdots, n$ denote historical returns, and μ_0 is a required mean rate of return. This approach replaces the traditional Markowitz use of variance as a measure of risk with a lower-tail expectation measure.

- The number of assets, p, is potentially quite large in these problems.
- Linear inequality constraints can easily be added to the problem to prohibit short sales, etc.
- Interior point methods are fine, but POGS may have advantages in larger problems.

Example 2: Smoothing the Quantile Regression Process

Let $\tau_1,\cdots,\tau_m \subset (0,1)$ denote an equally spaced grid and consider

$$\min_{\beta(\tau)\in\mathsf{R}^{\mathrm{mp}}}\{\sum_{k=1}^{\mathrm{m}}\sum_{i=1}^{n}w_{\tau_{k}}\rho_{\tau_{k}}(y_{i}-x_{i}^{\top}\beta(\tau_{k}))\mid\sum_{k}(\Delta^{2}\beta(\tau_{k}))^{2}\leqslant M\}.$$

Imposes a conventional L_2 roughness penalty on the quantile regression coefficients.

- Implemented recently in POGS by Shenoy, Gorinevsky and Boyd (2015) for forecasting load in a large power grid setting,
- Smoothing, or borrowing strength from adjacent quantiles, can be expected to improve performance,
- Many gory details of implementation remain to be studied.

Conclusions and Lingering Doubts

- Optimization can replace sorting
- Simplex is just steepest descent at successive vertices
- Log barriers revive Newton method for linear inequality constraints
- Proximal algorithms revive gradient methods
- Statistical vs computational accuracy?
- Quantile models as global likelihoods?
- Multivariate, IV, extensions?